

ON SPATIAL DYNAMICS*



Klaus Desmet

Universidad Carlos III de Madrid, Department of Economics, Calle Madrid 126, 28903 Getafe, Spain. E-mail: klaus.desmet@uc3m.es

Esteban Rossi-Hansberg

Princeton University, Department of Economics, 309 Fisher Hall, Princeton, NJ 08544. E-mail: erossi@princeton.edu

ABSTRACT. It has long been recognized that the forces that lead to the agglomeration of economic activity and to aggregate growth are similar. Unfortunately, few formal frameworks have been advanced to explore this link. We critically discuss the literature and present a simple framework that can circumvent some of the main obstacles we identify. We discuss the main characteristics of an equilibrium allocation in this dynamic spatial framework, present a numerical example to illustrate the forces at work, and provide some supporting empirical evidence.

1. INTRODUCTION

Economists have long discussed the relationship between agglomeration and growth. As Lucas (1988) points out, not only are both phenomena related to increasing (or constant) returns to scale, but in many contexts agglomeration forces are the source of the increasing returns that lead to growth. Krugman (1997), after providing a detailed overview of the different economic forces that can explain both phenomena, identifies probably the most important challenge of this literature: the difficulty of developing a common framework that incorporates both the spatial and the temporal dimensions. In other words, what is needed is a dynamic spatial theory. In this brief paper, we review the recent literature that has emerged to deal with some of the main links between growth and regional economics, discuss the problems that this literature faces, sketch a framework that we believe can be used to further explore the links between the spatial and temporal dimensions, and provide some empirical evidence consistent with the forces present in this framework.

The dynamics of the distribution of economic activity in space have been studied using three distinct approaches. A first family of models consists of

*Prepared for the conference “Regional Science: The Next Fifty Years,” on the occasion of the 50th anniversary of the *Journal of Regional Science*. We thank the editor, Gilles Duranton, and conference participants for useful comments. Financial support from the Comunidad de Madrid (PROCIUDAD-CM), the Spanish Ministry of Science (ECO2008-01300) and the European Commission (EFIGE grant 225343) is gratefully acknowledged.

Received: March 2009; revised: June 2009; accepted: June 2009.

dynamic extensions of *New Economic Geography* models. These models tend to have a small number of locations, typically two. Agglomeration is driven by standard Krugman (1991) pecuniary externalities operating through real wages. The models are usually made dynamic by adding innovation in product quality as in Grossman and Helpman (1991a,b). There is a wide variety of particular specifications, some of which include capital accumulation or other forms of innovation. Baldwin and Martin (2004) provide a nice survey of this literature. They highlight the possibility of “catastrophic” agglomeration, implying that only one region accumulates factors. More generally, agglomeration and innovation reinforce each other, creating growth poles and sinks. The emergence of regional imbalances is accompanied by faster aggregate growth and higher welfare in all regions.¹

The contribution of this first strand of the literature is important, as it enhances our understanding of the common forces underlying growth and agglomeration. However, the spatial predictions are rather limited. The focus on a small number of locations does not allow this literature to capture the richness of the observed distribution of economic activity across space, thus restricting the way these models are able to connect with the data. It advances statements about how unequal two regions are, but there is no sense in which one can have a hierarchy of agglomerated areas. One could of course try to generalize these models to more than a few regions. The problem is that the analytical tractability breaks down when one deals with more than two or three regions. Some progress could be made numerically, using dynamic extensions of continuous space *New Economic Geography* frameworks, like the one in Fujita, Krugman, and Venables (2001, Chapter 17), but little has been done so far. Therefore, these models remain mostly useful as analytical tools, rather than as guides to doing empirical work.

A second family of models aims to explain the distribution of city sizes. In general, this literature only models, if at all, space *within* cities, but not the location of cities *across* space. Early contributions include Black and Henderson (1999) and Eaton and Eckstein (1997). Black and Henderson (1999) propose a model of a dynamic economy with cities. Increasing returns in the form of externalities create cities and imply, apart from knife-edge parameter conditions, increasing returns at the aggregate level. Hence, as in the papers above, agglomeration leads to explosive growth. In contrast to the first strand of the literature, these theories have the advantage of explicitly modeling the cities in each location and allowing for heterogeneity in city characteristics. This comes at the cost of a black box agglomeration effect in the form of a production externality.

Within this second strand of the literature, the contribution of Gabaix (1999a) is key in establishing the link between the dynamic growth process

¹Readers interested in this strand of the literature should consult Baldwin and Martin (2004), Fujita and Thisse (2002, Chapter 11), and some of the specific papers, such as Baldwin, Martin, and Ottaviano (2001) and Martin and Ottaviano (1999, 2001).

of cities and the observed distribution of city sizes. He shows that Zipf's Law for cities—the size distribution is approximated by a Pareto distribution with coefficient one—can be explained by models that imply cities exhibiting scale-independent growth. For our purposes, the interesting part of this contribution is not so much the particular size distribution this growth process leads to, but rather the link it establishes between the dynamic growth process of particular production sites and the invariant distribution of economic activity in space. It is the growth process that leads to agglomeration (in the form of a size distribution with a fat right tail with many large cities), and not the other way around. Following Gabaix (1999a), many papers have been built on this basic insight, which had already been used in other applications in macroeconomics. Eeckhout (2004), for example, proposes a simple model in which cities grow by receiving scale-independent shocks, and uses the Central Limit Theorem to show that the resulting size distribution is log normal.

Gabaix (1999a,b) and Eeckhout (2004) postulate the growth rate of cities; they do not propose an economic theory of this growth process. The last generation of models in this second strand of the literature addresses this shortcoming by successfully establishing a link between economic characteristics that determine the growth process and economic agglomeration in cities. Duranton (2007) does so by proposing a growth process through the mobility of industries across cities as a result of innovations in particular sectors. Rossi-Hansberg and Wright (2007) also produce a particular city growth process as a result of adjustment in optimal city sizes and city entry. Córdoba (2008) discusses general properties that these models need to satisfy in order to yield a growth process consistent with particular characteristics of invariant distributions, like Zipf's Law. Some of these papers also establish a reverse link between the growth process and agglomeration. In Rossi-Hansberg and Wright (2007), for example, it is the organization of economic activity in cities that leads to the aggregate constant returns to scale necessary to generate balanced growth. In this sense, agglomeration of economic activity in a particular number and size of cities generates aggregate balanced growth.

The main limitation of the dynamic frameworks in this second strand of the literature is the lack of geography. Production happens in particular sites, but these sites are not ordered in space and the trade links between them are either frictionless or uniform. Cities are the units in which production is organized. The internal structure of cities is sometimes modeled as an area with land as a factor of production and agents facing transport and/or commuting costs. However, geography is only modeled *within* cities, not *across* them. In this sense, these models do not present dynamic spatial theories that can be contrasted to the observed distribution of economic activity in space.

The third strand of the dynamic spatial literature incorporates fully forward-looking agents and factor accumulation into models with a continuum

of geographically ordered locations.² It also allows for either capital mobility or some form of spillovers or diffusion between regions (Quah, 2002; Brito, 2004; Brock and Xepapadeas, 2008a,b; Boucekkine, Camacho, and Zou, 2009). Apart from these interactions, points in space are still completely isolated from each other. We review the particular structure of these problems in Section 3 below. For now it suffices to say that progress here has been mostly restricted to formulating the necessary and sufficient conditions for efficient allocations and, in some cases, the corresponding conditions characterizing rational expectation equilibria. Few substantive results have been advanced.

The remainder of this paper is organized as follows. In Section 2, we go further into the importance of developing spatial frameworks that can be compared with the data, some of the difficulties of doing this, and the comparison with trade frameworks, like that in Eaton and Kortum (1999). Section 3 discusses some of the setups with continuous space that have been analyzed for the case of forward-looking agents. Section 4 then proposes a simple endogenous growth spatial framework in which innovation decisions are optimally not forward-looking, and it uses a numerical example to shed light on the different forces present in this framework. Section 5 presents some basic evidence from the United States on the forces highlighted in Section 4, and Section 6 concludes.

2. THE IMPORTANCE OF SPACE

Incorporating geographically ordered space (or land) is important for two main reasons. Land at a particular location is a rival and nonreplicable input of production, and land is geographically ordered in a way that matters for economic activity. The latter claim has been documented extensively: patents cite geographically close-by patents (Jaffe, Trajtenberg, and Henderson, 1993), firms co-locate (Ellison and Glaeser, 1997; Duranton and Overman, 2005, 2008), and in general there is ample evidence of substantial trade costs, mobility costs, commuting costs, and other costs that increase with distance. The use of land as a nonreplicable input of production requires, perhaps, some additional explanation. Economic activity at a particular location is, of course, endogenous, so the factors operating at a given location can be replicated. Nevertheless, since land is an input of production, increasing factors at a given location leads to decreasing returns to scale and therefore dispersion.

It is obviously difficult to incorporate space into dynamic frameworks because it increases the dimensionality of the problem. Another difficulty of incorporating a continuum of locations in geographic space is that, in the presence of

²We discuss in more detail the importance of using a continuum of locations in the next section, but the evidence seems to suggest that the observed patterns are very different when land, and not only cities, is incorporated into the analysis. In particular, Holmes and Lee (2008) show that the distribution of employment across equal sized squares in space has a significantly lower tail than the one for cities. They also show that for space, and in contrast with cities, growth rates are not independent of scale (Gibrat's Law).

mobility frictions like transport or commuting costs, clearing factor and goods markets are not trivial. The reason is that how many goods or factors are lost in transit depends on mobility and trade patterns, which in turn depend on factor prices that are the result of market clearing. Hence, to impose market clearing it is necessary to know the number of goods lost in transit. That is, factor prices at each location depend on the equilibrium pattern of trade and mobility at all locations. This yields a problem that in many cases is intractable.

The trade literature has circumvented this difficulty by analyzing the case of a finite (though potentially large) number of locations in the presence of random realizations of productivity for a continuum of goods (see, e.g., Eaton and Kortum, 2002). In such a framework, the only relevant equilibrium variable is the share of exported and imported goods, which is well determined by the properties of the distributions of the maximum of the productivity realizations. This has proven to be an effective way to deal with this difficulty. However, it does not allow us to talk about trade in particular sectors, since only aggregate trade flows are determined in equilibrium. This is an important drawback if we want to study geography models that focus on spatial growth across industries. Since the empirical evidence shows that different sectors exhibit very different spatial growth patterns, this is a relevant issue (see, e.g., Desmet and Fafchamps, 2006; Desmet and Rossi-Hansberg, 2009a).

Another way of solving this problem is to clear markets sequentially. Suppose space is linear and compact. Then we can start at one end of the space interval and accumulate production minus consumption in a given market (properly discounted by transport or commuting costs) until we reach the end of the interval. At the boundary, “excess supply” has to be equal to zero in order for markets to clear. This method, proposed in Rossi-Hansberg (2005), is fairly easy to apply, but it can only be used in one-dimensional (or two-dimensional and symmetric) compact setups. Extending this formulation to nonsymmetric two-dimensional spatial setups (like reality!) is a theoretical challenge.

In Section, 4 we sketch a model that uses this form of market clearing. Our view is that it is possible to improve our understanding of dynamic spatial interactions using fully specified economic dynamic equilibrium models. In contrast, many geographers rely on so-called agent-based models to capture the complexities of spatial dynamics. The drawback of these models is that they lack economic fundamentals (see Irwin, 2009, in this volume on the use of agent-based models by economists).

3. SPATIAL MODELS WITH FORWARD-LOOKING AGENTS

The few papers that have studied a fully dynamic setup with a continuum of locations normally focus on the problem of a planner who allocates resources. We present two examples below. Spatial interactions are introduced in two different ways: a first one by allowing for capital mobility, and a second one by assuming a spatial capital externality. Neither of them introduces land as an

input of production, although given that technology is not necessarily assumed to be constant returns to scale, it could be easily incorporated through absentee landlords.

The spatial setup is the real line and time is continuous. Let $c(\ell, t)$ denote consumption, $L(\ell, t)$ population, and $k(\ell, t)$ capital at location ℓ and time t . A central planner then maximizes the sum of utilities of all agents, all of whom discount time at rate β . Production requires only capital, $k(\ell, t)$, which depreciates at rate δ . Total factor productivity is given by $Z(\ell, t)$. The change in capital at a particular location is therefore equal to production minus depreciation minus consumption plus the capital received from other locations. Boucekkine et al. (2009) show how this last term can be expressed as the second partial derivative of capital across locations: essentially, it is just the difference between the flow of capital from the regions to the left minus the flow of capital flowing to the regions to the right.³ This law of motion of capital, a parabolic differential equation, and in particular the spatial component entering through the second order term, introduces space into the problem. In addition, capital at all locations at time 0 is assumed to be known, and since the real line is infinite, a transversality condition on capital is required. Hence, the problem solved by Boucekkine et al. (2009) becomes:

$$\begin{aligned} & \max_c \int_0^\infty \int_{\mathbb{R}} U(c(\ell, t)) L(\ell, t) e^{-\beta t} d\ell dt \\ & \text{subject to} \\ & \frac{\partial k(\ell, t)}{\partial t} - \frac{\partial^2 k(\ell, t)}{\partial \ell^2} = Z(\ell, t) f(k(\ell, t)) - \delta k(\ell, t) - c(\ell, t) \\ & k(\ell, 0) = k_0(\ell) > 0 \\ & \lim_{\ell \rightarrow \pm\infty} \frac{\partial k(\ell, t)}{\partial \ell} = 0. \end{aligned}$$

Brock and Xepapadeas (2008b) and Brito (2004) solve similar problems, but with different preferences. In fact, Boucekkine et al. (2009) show that for general preferences this is an “ill-posed” problem in the sense that the initial value of the co-state does not determine its whole dynamic path. This is a general problem in spatial setups. One can address this issue either by considering particular solutions (like the type of cyclical perturbation analysis found in many studies) or by putting strong restrictions on preferences. Boucekkine et al. (2009) show that some progress can be made by focusing on the linear case.

³If a region is an interval in space, the capital received from other regions is the difference in the partial derivatives of capital at the two boundary points. When in the limit a region becomes a point in space, the difference in these partial derivatives equals the second partial derivative.

Brock and Xepapadeas (2008b) study a similar problem in a compact interval R , given by

$$\max_c \int_0^\infty \int_R U(k(\ell, t), c(\ell, t), X(\ell, t)) L(\ell, t) e^{-\beta t} d\ell dt$$

subject to

$$\frac{\partial k(\ell, t)}{\partial t} = f(k(\ell, t), c(\ell, t), X(\ell, t))$$

$$X(\ell, t) = \int_{\ell \in R} \omega(\ell - \ell') k(\ell', t) d\ell'$$

$$k(\ell, t) = k_0(\ell) > 0,$$

where $X(\ell, t)$ is an externality that affects production and utility, and f now refers to production minus consumption plus an additional term reflecting the direct effect of the externality on the law of motion of capital. In contrast to the problem of Boucekkine et al. (2009), there is no capital mobility, which eliminates a huge difficulty. Instead, the spatial component is introduced through the externality, which is just a kernel of capital at all locations. This is an interesting problem, since it incorporates diffusion, although not mobility. As in the previous case, the authors can derive the Pontryagin necessary conditions for an optimum and, under more restrictive assumptions, sufficient conditions. Solving for stable steady states remains, nevertheless, an exercise of finding whether or not spatially uniform steady states are stable. In other words, they are unable to fully analyze spatially nonuniform steady states. This is progress, although it does not amount to a complete analysis of the problem.

The lack of a complete solution to the problems above is hardly the fault of the authors working on them. These problems are complicated and, absent more structure, it is hard to extract general insights. The main problem seems to be that agents are forward-looking and thus need to understand the whole future path to make current decisions. Modeling space implies understanding the whole distribution of economic activity over space and time for each feasible action. One way around this difficulty is to impose enough structure—either on the diffusion of technology or on the mobility of agents and land ownership—so that agents do not care to take the future allocation paths into account, given that they are out of their control and do not affect the returns from current decisions. In the next section, we present an example of such a framework.

4. AN ALTERNATIVE MODEL WITH FACTOR MOBILITY AND DIFFUSION

In Desmet and Rossi-Hansberg (2009b), we introduce a model in which locations accumulate technology by investing in innovation in one of two industries and by receiving spillovers from other locations. The key to making such a rich structure computable is that diffusion, together with labor mobility and diversified land ownership, implies that the decisions of where to locate

and how much to invest in technology do not depend on future variables. As a result, in spite of being forward-looking, agents solve static problems. The dynamics generated by the model lead to locations changing occupations and employment density continuously, but in the aggregate the economy converges on average to a balanced growth path.

Desmet and Rossi-Hansberg (2009b) study an economy with two sectors and analyze the sectoral interaction in generating innovation. They use the model to explain the observed evolution in the spatial distribution of economic activity in the United States. To give a sense of the forces at work in that model, we here present a simpler version of the setup with only one good (and therefore no specialization decision or cross-industry innovation effects). In this version of the model, factor mobility is frictionless, and trade is just the result of agents holding a diversified portfolio of land across locations.

Land is given by the unit interval $[0, 1]$, time is discrete, and total population is \bar{L} . We divide space into “counties” (connected intervals in $[0, 1]$), each of which has a local government. Agents solve

$$\max_{\{c(\ell, t)\}_{t=0}^{\infty}} E \sum_{t=0}^{\infty} \beta^t U(c(\ell, t))$$

subject to

$$w(\ell, t) + \frac{\bar{R}(t)}{\bar{L}} = p(\ell, t) c(\ell, t) \text{ for all } t \text{ and } \ell,$$

where $p(\ell, t)$ is the price of the consumption good and $w(\ell, t)$ denotes the wage at location ℓ and time t . Total land rents per unit of land at time t are denoted by $\bar{R}(t)$, so that $\bar{R}(t)/\bar{L}$ is the dividend from land ownership received by agents, assuming that agents hold a diversified portfolio of land in all locations. Free mobility implies that utilities equalize across regions each period.

The inputs of production are land and labor. Production per unit of land is given by

$$x(L(\ell, t)) = Z(\ell, t) L(\ell, t)^\mu,$$

where $\mu < 1$, $Z(\ell, t)$ denotes TFP, and $L(\ell, t)$ is the amount of labor per unit of land used at location ℓ and time t . The problem of a firm at location ℓ is thus given by

$$\max_{L(\ell, t)} (1 - \tau(\ell, t)) (p(\ell, t) Z(\ell, t) L(\ell, t)^\mu - w(\ell, t) L(\ell, t)),$$

where $\tau(\ell, t)$ denotes taxes on profits charged by the county government.

The government of a county can decide to buy an opportunity to innovate by taxing local firms $\tau(\ell, t)$. In particular, a county can buy a probability $\phi \leq 1$ of innovating at a cost $\psi(\phi)$ per unit of land. This cost $\psi(\phi)$ is increasing and convex in ϕ , and proportional to wages. If a county innovates, all firms in the county have access to the new technology. A county that obtains the chance

to innovate draws a technology multiplier $z(\ell)$ from a Pareto distribution with lower bound 1,⁴ leading to an improved level of TFP, $z_\ell Z_i(\ell, t)$, where

$$\Pr[z < z_\ell] = \left(\frac{1}{z}\right)^a.$$

The risk-neutral government of county G , with land measure I , will then maximize

$$(1) \quad \max_{\phi(\ell, t)} \int_G \frac{\phi(\ell, t)}{a-1} p(\ell, t) Z(\ell, t) L(\ell, t)^\mu d\ell - I\psi(\phi).$$

The benefits of the extra production last only one period. Since a county is by assumption small and innovation diffuses geographically, a county's innovation decision today does not affect its expected level of technology tomorrow. This implies that governments need not be forward-looking when choosing the optimal level of investment in innovation. Note the scale effect in the previous equation: high employment density locations will optimally innovate more (and so will high-price and high-productivity locations). This is consistent with the evidence presented by Carlino, Chatterjee, and Hunt (2007). They show that a doubling of employment density leads to a 20 percent increase in patents per capita.

The timing of the problem is key. Innovation diffuses spatially between time periods.⁵ So, before the innovation decision, location ℓ has access to

$$Z_i(\ell, t+1) = \max_{r \in [0,1]} e^{-\delta|\ell-r|} Z(r, t)$$

which of course includes its own technology. This means that in a given period each location has access to the best spatially discounted technology of the previous period. Agents then costlessly relocate, ensuring that utility is the same across all locations. After labor moves, counties invest in innovation. Assuming wages are set before the innovation decision, the fact that agents hold a diversified portfolio of land in all locations implies that they need not be forward-looking when deciding where to locate. Note also that by holding a diversified portfolio of land, rents are redistributed from high-productivity to low-productivity locations. As a result, high-productivity locations run trade surpluses, and low-productivity locations run trade deficits.

In addition to the geographic diffusion of innovations, transport costs are another source of agglomeration. For simplicity we assume iceberg transport costs, so if one unit of the good is transported from ℓ to r , only $e^{-\kappa|\ell-r|}$ units of the good arrive in r . Hence, if goods are produced in ℓ and consumed in r , $p(r, t) = e^{\kappa|\ell-r|} p(\ell, t)$. As described in Section 2, goods markets clear sequentially. The

⁴Using the Pareto distribution simplifies the analytics, but is not essential to the argument.

⁵For early work on the spatial diffusion of technologies, see Griliches (1957) and Hågerstrand (1967).

stock of excess supply between locations 0 and ℓ , $H(\ell, t)$, is defined by $H(0, t) = 0$ and by the differential equation

$$\frac{\partial H(\ell, t)}{\partial \ell} = \theta(\ell, t)x(\ell, t) - c(\ell, t) \left(\sum_i \theta(\ell, t)L(\ell, t) \right) - \kappa |H(\ell, t)|.$$

At each location the change in the excess supply is the difference between the quantity produced and the quantity consumed, net of the shipping cost in terms of goods lost in transit. Then, the goods market clears if $H(1, t) = 0$. The labor market clearing condition is given by

$$\int_0^1 L(\ell, t) d\ell = \bar{L}, \text{ all } t.$$

Computing an equilibrium of this economy is clearly feasible. Given initial productivity functions, we can solve for production in all locations, for the wages that equalize utility and clear the national labor market, for the prices that clear the goods market, and for the resulting average land rents, which are added to agents' income. This determines the location of agents and the investments in innovation. After productivity is realized, we compute actual production, actual distributed land rents, and trade. Overnight there is diffusion, which determines the new productivity function. Since decisions are based on current outcomes only, computing an equilibrium involves solving a functional fixed point each period, but it does not involve calculating rational expectations.

What can we learn from this model? Although the model is extremely simple, it has two forces that are interesting when thinking about spatial dynamics. On the one hand, although technology is constant returns in land and labor, it exhibits local decreasing returns to labor, because locally land cannot be replicated. This is a form of local congestion that spreads employment across space given identical technology levels. On the other hand, agglomeration is the result of the diffusion of technology. Areas with high levels of employment innovate more, since the incentives to innovate are larger there. Since diffusion decreases with distance, areas close to high-employment clusters become high-productivity areas. This attracts employment and leads to more innovation. As usual, the balance between the congestion and agglomeration forces determines the spatial landscape.

The same forces that lead to particular spatial employment patterns also explain aggregate growth. Dispersion implies more uniform, but smaller, incentives to innovate. In contrast, concentration implies that fewer locations innovate, but each of them innovates more. More diffusion implies that the second (extensive) effect is less important and that aggregate growth is generally greater.

Perhaps surprisingly, higher trade costs imply more concentrated production, which in turn may lead to more growth. Although higher trade costs imply static efficiency losses, they also lead to dynamic gains through increased

concentration and innovation, an effect reminiscent of the one in Fujita and Thisse (2003). A clear empirical implication emerges from the theory: more concentration of employment in surrounding areas leads to higher innovation and growth. This effect is the result of two forces. First, more concentration as a result of, say, transport costs, leads to more innovation. Second, more innovation in certain areas leads, through diffusion, to productivity growth in neighboring areas (see, e.g., Rosenthal and Strange, 2008, for evidence on this mechanism).⁶

The model presented above has only one industry, so by construction it is not suited to study cross-industry effects. In Desmet and Rossi-Hansberg (2009b), we present a version of the model with two industries. In that case, another spatial link between the distribution of economic activity and growth emerges. Locations near clusters of firms in one sector, say, manufacturing, experience high prices of the other good, say, services, since their proximity to manufacturing locations allows them to sell services paying small trade costs. This channel works through trade: neighboring areas that are specialized in manufacturing will import services, thus pushing up the relative price of services. As a result, locations close to manufacturing clusters tend to have high employment and high prices in services and therefore will tend to innovate in services. Hence, being near clusters in the other industry is also a source of growth and innovation. However, note that this force operates through imports, whereas the diffusion force operates through employment. In the next section, we present some evidence supporting these predictions.

Figure 1 presents a numerical simulation from the framework with two sectors, manufacturing and services. The model used to compute the figure is identical to the one presented in Desmet and Rossi-Hansberg (2009b). The calibration we use here to illustrate the outcome of the model sets $\delta = 50$, $\kappa = 0.005$, the elasticity of substitution between manufactures and services equal to 0.5 and the Pareto parameter $\alpha = 35$. We use a cost function given by $\psi(\phi) = \psi_1 + \psi_2/(1 - \phi)$ and set $\psi_2 = -\psi_1 = 0.003851w(\ell, t)$. Desmet and Rossi-Hansberg (2009b) provides a careful discussion on the effect of most parameters on equilibrium allocations. The figure shows a contour map of productivity in time and space. Space is the unit interval (divided in 500 locations), and we run the model for 100 periods. We added a scale that shows how different levels of productivity are represented. The figure helps to identify where clusters are located and how they are created and destroyed over time.

We use initial conditions that imply that locations close to the upper bound are good in manufacturing, whereas all locations have an initial productivity in services equal to 1. These initial conditions imply that manufacturing starts innovating first and only in the upper regions. As we argued, diffusion implies that regions that innovate are clustered. As a result, productivity growth

⁶Duranton and Overman (2005, 2008) present detailed and strong evidence of co-location in the U.K. Their focus is on regional agglomeration mechanisms within and between industries. Unfortunately, they do not directly address the link between growth and regional agglomeration.

DESMET AND ROSSI-HANSBERG: ON SPATIAL DYNAMICS

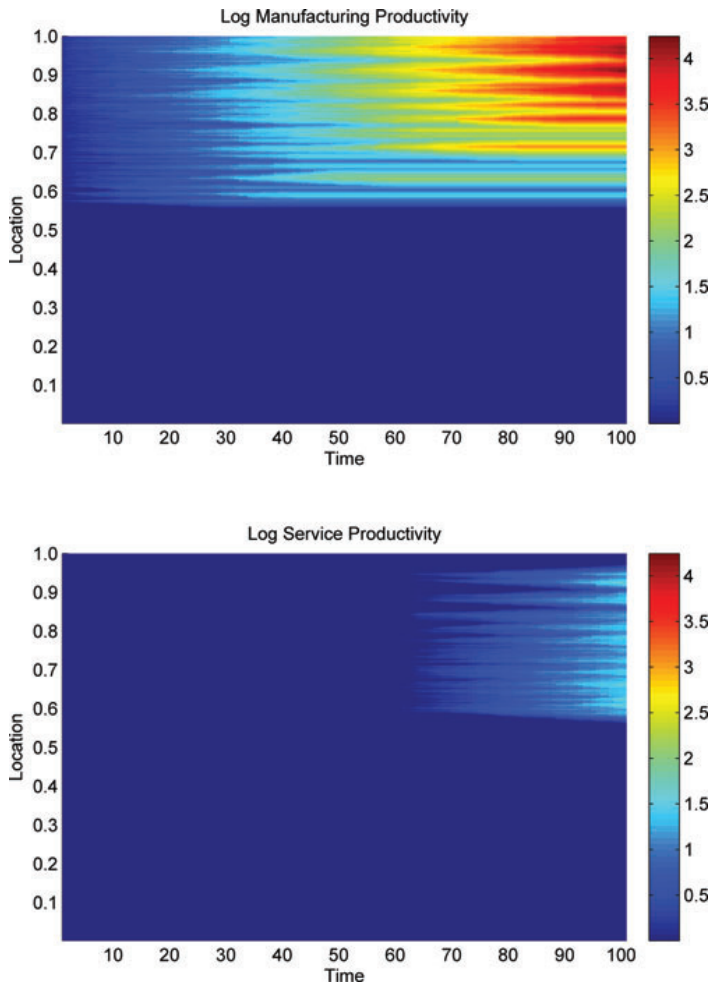


FIGURE 1: An Example.

happens in concentrated areas. This is an expression of the first effect discussed above. Given that innovation clusters coincide with employment clusters, the model is able to generate “spikes” of economic activity, which could be interpreted as cities. The result is then similar to models of systems of cities (see, e.g., Black and Henderson, 1999), but with the advantage that in our framework space is a continuum.

In period 63 some scattered service areas, which are close to manufacturing clusters, start innovating. This innovation happens in clusters too and, more important, next to manufacturing areas. Relative prices of services are high next to clusters of manufacturing production as a result of transport costs and trade. This leads to endogenously higher employment and more innovation in services. This is an expression of the second effect discussed above.

It is important to understand how productivity growth in the service sector gets jump-started. Assuming an elasticity of substitution less than one, the sector with the higher relative productivity growth loses employment share. Initially, when only manufacturing is innovating, the share of employment in services is gradually increasing. Since gains from innovation in a given sector depend on employment in that sector, at some point the service sector becomes large enough, allowing for innovation to take off. This mechanism provides an endogenous stabilization mechanism that tends to increase the productivity of one of the sectors when the economy experiences fast productivity growth in the other sector. The result is that by period 100 both sectors are growing at a roughly constant rate of around 3 percent.

In Desmet and Rossi-Hansberg (2009b) we match the model to some of the main features of the U.S. economy over the last 25 years. Doing so allows us to analyze the effect of changes in certain relevant parameter values. As mentioned before, we show, for example, that higher transportation costs may yield dynamic welfare gains through increased spatial concentration leading to more innovation. As argued by Holmes (2009) in this volume, having a fully specified theoretical model that can be matched to the data and run on the computer has much to offer to the field of regional and urban economics. Most of the empirical work in the field has taken a reduced-form, rather than a structural, approach.⁷

5. SOME EMPIRICAL EVIDENCE

The model in Section 4 illustrates two main forces that mediate spatial dynamics. The first one is a “spillover” effect by which locations close to other locations in the same sector grow faster because they benefit from innovation investments close-by. The second is a “trade” effect by which locations close to areas that import a particular good experience high prices for that good, thus providing incentives to innovate in that sector. If these effects are the cornerstone of spatial dynamics, as the model above postulates, we should be able to find them in the data.

Using U.S. county data for the period 1980–2000 from the Bureau of Economic Analysis, we first construct two kernels to measure the importance of the “spillover” and the “trade” effect. For each county, the first kernel sums employment over all other counties, exponentially discounted by distance. To

⁷For a recent example of this structural approach in regional economics, see Holmes (2008).

compute the second kernel, we first measure county imports in a particular sector as the difference between the county's consumption and production in that sector.⁸ For each county, the second kernel then sums sectoral imports over all counties, exponentially discounted by distance.⁹ This constitutes a measure of the excess demand experienced by a county in a particular sector. With these two kernels in hand, we run the following regression:

$$\begin{aligned} \log Emp_{\ell}^i(t+1) - \log Emp_{\ell}^i(t) = & \alpha + \beta_1 \log Emp_{\ell}^i(t) + \beta_2 \log (EK_{\ell}^i(t)) \\ & + \beta_3 \log (IK_{\ell}^i(t)), \end{aligned}$$

where $Emp_{\ell}^i(t)$ denotes employment, $EK_{\ell}^i(t)$ the employment kernel, and $IK_{\ell}^i(t)$ the imports kernel, for sector i , county ℓ , and period t .^{10,11}

Table 1 presents the results for different discount rates. We fix the discount rate for the employment kernel at 0.1 (implying the effect declines by half every 7 km),¹² and let the decay parameter for the import kernel vary between 0.07 and 0.14 (implying the effect declines by half every 5 to 10 km).¹³ We present four sets of regressions, the first two present the results for the service sector for the decades 2000–1990 and 1990–1980, and the last two present the same regressions for the industrial sector (manufacturing plus construction).

To illustrate our results, focus on the case of a decay parameter in the import kernel of 0.1 (identical to the one in the employment kernel). In services, we find that for the 1990s a 1 percent increase in the initial employment kernel leads to a 0.006 percent increase in county service employment between 1990 and 2000. The coefficient on the employment kernel does not change much across different decay parameters and across both sectors. We obtain a different result for the 1980–1990 decade, where the coefficients are still positive and significant, but the coefficient in industry is substantially larger.¹⁴

⁸A county's consumption in a given sector is obtained by multiplying the national share of earnings in that sector by the county's total earnings. A county's production in a given sector is simply measured by its earnings in that sector. Note that this calculation does not take into account international trade, most of which is in goods. However, since this changes the level of imports in a similar way in all counties, it should not affect our calculations significantly.

⁹Note that, according to the theory, the discount rate should be related to transport costs.

¹⁰Since the import kernel measures a discounted sum of imports in a given sector, this measure may be positive or negative. We can therefore not simply take the natural logarithm. In the regression, we use the natural logarithm of the kernel when the kernel is positive and minus the natural logarithm of the absolute value of the kernel when it is negative.

¹¹Since we include the log of employment in county ℓ as a separate regressor, the employment kernel does not include employment in county ℓ . In contrast, the import kernel does include imports by county ℓ .

¹²This sharp geographic decline in spillovers is consistent with findings in Rosenthal and Strange (2008), who report that human capital spillovers within a range of 5 miles are four to five times larger than at a distance of 5 to 25 miles.

¹³Using detailed micro-data, Hillberry and Hummels (2008) document that the value of shipments within the same 5-digit zip code are three times higher than those outside the zip code.

¹⁴Dumais, Ellison, and Glaeser (2002) provide firm-level evidence of a "spillover" effect in the manufacturing industry. For a detailed discussion of the effect of current employment on sectoral growth rates, see Desmet and Rossi-Hansberg (2009a).

TABLE 1: The Effect of Employment and Import Kernels on U.S. Employment Growth Rates

Decay Emp. Kernel:		0.1 (half life 7 km)									
Decay Imp. Kernel:		0.07	0.08	0.09	0.10	0.11	0.12	0.13		0.14	
Half-Life Imp. Kernel (km):		9.9	8.7	7.7	6.9	6.3	5.8	5.3		5.0	
Dependent variable: Log(Service Employment 2000)-Log(Service Employment 1990)											
Log(Serv. Emp. 1990)		0.00346 [1.29]	0.00383 [1.43]	0.00409 [1.53]	0.00426 [1.59]	0.0043 [1.61]	0.00443 [1.66]*	0.00451 [1.69]*		0.00458 [1.72]*	
Log(Serv. Emp. Kernel 1990)		0.00624 [4.55]***	0.00603 [4.40]***	0.00587 [4.28]***	0.00576 [4.20]***	0.00572 [4.17]***	0.00563 [4.11]***	0.00557 [4.07]***		0.00552 [4.03]***	
Log(Serv. Imp. Kernel 1990)		-0.00028 [0.74]	0.00014 [0.36]	0.00044 [1.15]	0.00065 [1.67]*	0.00073 [1.85]*	0.00089 [2.26]**	0.00101 [2.55]**		0.00113 [2.87]***	
Constant		0.18715 [8.08]***	0.18406 [7.95]***	0.18195 [7.86]***	0.1805 [7.80]***	0.18018 [7.79]***	0.17918 [7.75]***	0.17855 [7.73]***		0.17789 [7.71]***	
Observations		2277	2277	2277	2277	2277	2277	2277		2277	
R-squared		0.0131	0.013	0.0135	0.0141	0.0144	0.0151	0.0157		0.0165	
Dependent variable: Log(Service Employment 1990)-Log(Service Employment 1980)											
Log(Serv. Emp. 1980)		0.04007 [13.89]***	0.04012 [13.91]***	0.04028 [13.99]***	0.04034 [14.02]***	0.04026 [14.00]***	0.04024 [14.00]***	0.04023 [14.00]***		0.0402 [13.99]***	
Log(Serv. Emp. Kernel 1980)		0.01013 [6.89]***	0.01003 [6.82]***	0.00987 [6.72]***	0.00977 [6.67]***	0.00977 [6.67]***	0.00973 [6.65]***	0.00975 [6.66]***		0.00978 [6.68]***	
Log(Serv. Imp. Kernel 1980)		0.00153 [3.72]***	0.00176 [4.22]***	0.00208 [4.96]***	0.00232 [5.49]***	0.00236 [5.58]***	0.00249 [5.86]***	0.00251 [5.90]***		0.00245 [5.76]***	
Constant		-0.19616 [8.12]***	-0.19598 [8.11]***	-0.19655 [8.15]***	-0.19644 [8.15]***	-0.19573 [8.12]***	-0.19524 [8.11]***	-0.19523 [8.11]***		-0.19522 [8.11]***	
Observations		2616	2616	2616	2616	2616	2616	2616		2616	
R-squared		0.1191	0.1204	0.1227	0.1245	0.1248	0.1259	0.1261		0.1255	

Continued

TABLE 1: Continued

Decay Emp. Kernel:		0.1 (half life 7 km)									
Decay Imp. Kernel:		0.07	0.08	0.09	0.10	0.11	0.12	0.13	0.14		
Half-Life Imp. Kernel (km):		9.9	8.7	7.7	6.9	6.3	5.8	5.3	5.0		
Dependent variable: $\text{Log}(\text{Industry Employment } 2000) \cdot \text{Log}(\text{Industry Employment } 1990)$											
Log(Ind. Emp. 1990)		-0.053 [12.70]***	-0.0526 [12.59]***	-0.05222 [12.49]***	-0.05191 [12.41]***	-0.05166 [12.34]***	-0.05162 [12.32]***	-0.0514 [12.26]***	-0.05125 [12.21]***		
Log(Ind. Emp. Kernel 1990)		0.00624 [2.08]**	0.00617 [2.05]**	0.00607 [2.02]**	0.00602 [2.00]**	0.00605 [2.01]**	0.00592 [1.97]**	0.00584 [1.94]*	0.00567 [1.88]*		
Log(Ind. Imp. Kernel 1990)		0.00626 [10.71]***	0.00624 [10.54]***	0.00628 [10.49]***	0.00638 [10.59]***	0.00639 [10.53]***	0.00632 [10.36]***	0.00633 [10.36]***	0.00626 [10.22]***		
Constant		0.56213 [19.06]***	0.55913 [18.92]***	0.55678 [18.82]***	0.55473 [18.74]***	0.55276 [18.65]***	0.55297 [18.64]***	0.55164 [18.59]***	0.55117 [18.55]***		
Observations		2543	2543	2543	2543	2543	2543	2543	2543		
R-squared		0.1221	0.121	0.1206	0.1213	0.1209	0.1197	0.1197	0.1187		
Dependent variable: $\text{Log}(\text{Industry Employment } 1990) \cdot \text{Log}(\text{Industry Employment } 1980)$											
Log(Ind. Emp. 1980)		-0.03445 [7.35]***	-0.03389 [7.22]***	-0.03371 [7.18]***	-0.03338 [7.10]***	-0.03287 [6.99]***	-0.03256 [6.92]***	-0.0323 [6.86]***	-0.03207 [6.81]***		
Log(Ind. Emp. Kernel 1980)		0.03753 [10.77]***	0.03786 [10.89]***	0.03795 [10.92]***	0.03806 [10.97]***	0.0382 [11.03]***	0.0382 [11.05]***	0.03824 [11.08]***	0.03828 [11.10]***		
Log(Ind. Imp. Kernel 1980)		0.00176 [2.45]**	0.00223 [3.06]***	0.0024 [3.26]***	0.00265 [3.59]***	0.00297 [4.00]***	0.00313 [4.21]***	0.00332 [4.47]***	0.00348 [4.68]***		
Constant		0.1047 [3.21]***	0.09882 [3.02]***	0.09711 [2.97]***	0.09398 [2.88]***	0.08948 [2.73]***	0.0871 [2.66]***	0.08495 [2.60]***	0.08312 [2.54]**		
Observations		2857	2857	2857	2857	2857	2857	2857	2857		
R-squared		0.0417	0.0428	0.0432	0.044	0.045	0.0456	0.0463	0.047		

Absolute value of t -statistics in brackets.

* significant at 10 percent; ** significant at 5 percent; *** significant at 1 percent.

TABLE 2: The Effect of Employment and Import Kernels on U.S. Earnings Growth Rates

Decay Emp. Kernel:		0.1 (half-life 7 km)									
Decay Imp. Kernel:		0.07	0.08	0.09	0.10	0.11	0.12	0.13	0.14		
Half-Life Imp. Kernel (km):		9.9	8.7	7.7	6.9	6.3	5.8	5.3	5.0		
<i>Dependent variable: Log(Service Earnings 2000)-Log(Service Earnings 1990)</i>											
Log(Serv. Emp. Kernel 1990)		[12.86]*** 0.04436	[12.98]*** 0.04472	[13.06]*** 0.04494	[13.14]*** 0.04517	[13.15]*** 0.04518	[13.18]*** 0.04528	[13.20]*** 0.04532	[13.22]*** 0.04536		
Log(Serv. Emp. Kernel 1990)		[5.15]*** 0.00964	[4.97]*** 0.00929	[4.84]*** 0.00905	[4.73]*** 0.00882	[4.68]*** 0.00874	[4.62]*** 0.00862	[4.58]*** 0.00854	[4.54]*** 0.00847		
Log(Serv. Imp. Kernel 1990)		[1.22]*** 0.00062	[2.27]*** 0.00117	[3.01]*** 0.00156	[3.70]*** 0.00193	[3.89]*** 0.00204	[4.21]*** 0.00222	[4.48]*** 0.00237	[4.73]*** 0.00251		
Constant		[2.84]*** 0.08251	[2.77]*** 0.08013	[2.72]*** 0.07879	[2.67]*** 0.07725	[2.68]*** 0.07742	[2.66]*** 0.07682	[2.66]*** 0.07669	[2.65]*** 0.0765		
Observations		2745	2745	2745	2745	2745	2745	2745	2745		
R-squared		0.0911	0.0923	0.0936	0.0951	0.0956	0.0965	0.0972	0.098		
<i>Dependent variable: Log(Service Earnings 1990)-Log(Service Earnings 1980)</i>											
Log(Serv. Emp. Kernel 1980)		[20.86]*** 0.07938	[20.91]*** 0.07944	[20.99]*** 0.07965	[21.03]*** 0.0797	[21.02]*** 0.07957	[21.01]*** 0.07947	[21.00]*** 0.07941	[20.97]*** 0.07932		
Log(Serv. Emp. Kernel 1980)		[7.12]*** 0.01383	[7.02]*** 0.01362	[6.92]*** 0.0134	[6.85]*** 0.01325	[6.83]*** 0.0132	[6.81]*** 0.01315	[6.82]*** 0.01316	[6.84]*** 0.01322		
Log(Serv. Imp. Kernel 1980)		[4.43]*** 0.00241	[5.19]*** 0.00285	[5.89]*** 0.00326	[6.47]*** 0.0036	[6.73]*** 0.00375	[6.98]*** 0.0039	[7.04]*** 0.00394	[6.88]*** 0.00386		
Constant		[5.65]*** -0.17965	[5.64]*** -0.17902	[5.67]*** -0.17955	[5.66]*** -0.17907	[5.61]*** -0.17744	[5.57]*** -0.17611	[5.56]*** -0.17554	[5.54]*** -0.17503		
Observations		2647	2647	2647	2647	2647	2647	2647	2647		
R-squared		0.2021	0.2043	0.2066	0.2087	0.2097	0.2107	0.211	0.2103		

Continued

TABLE 2: Continued

Decay Emp. Kernel:		0.1 (half life 7 km)									
Decay Imp. Kernel:		0.07	0.08	0.09	0.10	0.11	0.12	0.13	0.14		
Half-Life Imp. Kernel (km):		9.9	8.7	7.7	6.9	6.3	5.8	5.3	5.0		
Dependent variable: <i>Log(Industry Earnings 2000)-Log(Industry Earnings 1990)</i>											
Log(Ind. Emp. 1990)		0.02312 [3.32]***	0.02384 [3.42]***	0.02453 [3.52]***	0.02486 [3.56]***	0.02524 [3.62]***	0.02546 [3.65]***	0.02584 [3.70]***	0.02613 [3.74]***		
Log(Ind. Emp. Kernel 1990)		0.02212 [4.36]***	0.02232 [4.41]***	0.0225 [4.45]***	0.02239 [4.43]***	0.02251 [4.45]***	0.02247 [4.44]***	0.02252 [4.46]***	0.02242 [4.44]***		
Log(Ind. Imp. Kernel 1990)		0.00607 [6.15]***	0.0064 [6.41]***	0.00673 [6.67]***	0.00678 [6.69]***	0.00688 [6.72]***	0.00691 [6.72]***	0.00704 [6.83]***	0.00704 [6.82]***		
Constant		0.08837 [1.85]*	0.08216 [1.72]*	0.0765 [1.60]	0.07448 [1.55]	0.07109 [1.48]	0.06955 [1.45]	0.06641 [1.38]	0.06462 [1.34]		
Observations		2752	2752	2752	2752	2752	2752	2752	2752		
R-squared		0.0271	0.0282	0.0295	0.0295	0.0297	0.0297	0.0302	0.0301		
Dependent variable: <i>Log(Industry Earnings 1990)-Log(Industry Earnings 1980)</i>											
Log(Ind. Emp. 1980)		-0.02419 [3.89]***	-0.02376 [3.81]***	-0.02347 [3.76]***	-0.02324 [3.72]***	-0.02272 [3.63]***	-0.02231 [3.55]***	-0.02196 [3.49]***	-0.02165 [3.44]***		
Log(Ind. Emp. Kernel 1980)		0.05805 [12.87]***	0.05818 [12.91]***	0.05834 [12.96]***	0.05837 [12.97]***	0.05852 [13.02]***	0.0586 [13.06]***	0.05868 [13.09]***	0.05875 [13.11]***		
Log(Ind. Imp. Kernel 1980)		0.00168 [1.84]*	0.00191 [2.07]**	0.00213 [2.28]**	0.00223 [2.37]**	0.00249 [2.64]***	0.00269 [2.83]***	0.00287 [3.01]***	0.00302 [3.17]***		
Constant		0.32398 [7.36]***	0.31993 [7.23]***	0.31693 [7.15]***	0.31486 [7.09]***	0.31006 [6.96]***	0.30633 [6.86]***	0.30321 [6.78]***	0.30033 [6.70]***		
Observations		2758	2758	2758	2758	2758	2758	2758	2758		
R-squared		0.0587	0.059	0.0593	0.0595	0.0599	0.0603	0.0607	0.061		

 Absolute value of t -statistics in brackets.

* significant at 10 percent; ** significant at 5 percent; *** significant at 1 percent.

We also find a positive and robust “trade” effect. In 1980–1990, the effect seems to be similar in both industries. A 1 percent increase in the import kernel implies roughly a 0.002 percent increase in employment growth over the decade. In the 1990s, the effect is larger in industry and smaller in services. In almost all specifications, the “trade” effect is positive and significant. However, note that the model above leaves out another potential effect, namely, the growth effect of easier access to inputs in the same industry. This effect would imply, on its own, negative coefficients on the import kernel. The only case in which we obtain such a negative coefficient is when we use a low spatial discounting coefficient for the import kernel of services in 1990–2000. Since in that case the negative coefficient is statistically insignificant, we conclude that the trade effect seems to dominate the growth effects from easier access to inputs. However, more work is needed to explore these different effects.

Table 2 presents regressions similar to the ones in Table 1, but we now take sectoral earnings growth as the dependent variable. The results are similar, and, if anything, the coefficients are larger than for employment growth. According to the theory this should be the case, since the productivity and employment effect on innovation are complementary, as are the price and employment effects (see Equation (1)). As before, for virtually all decay parameters we find positive and significant “spillover” and “trade” effects.

6. CONCLUSION

In this paper, we have discussed the theoretical problems involved in the study of spatial dynamics. The literature consists of a set of frameworks that have only been partially understood and analyzed. To deal with some of the main obstacles in this literature, we have presented a simple framework that uses two main “tricks”: we make the required assumptions to make decisions static and we clear markets sequentially. This approach allowed us to underscore two key links between space and time, for which we have provided empirical support. In particular, we have shown that both the “spillover” and the “trade” innovation effects seem to be present in U.S. county data.

Undoubtedly, much work is still needed. First, we need to understand the basic frameworks better. In particular, we need to extract a set of robust insights from a model rich enough to be compared with the data. This requires a model with many locations and a distribution of economic activity varied enough to calculate standard statistics. Having two or three regions without land markets is not enough. Second, we need better ways of comparing these statistics with the data. What are the main attributes of the evolution of the distribution of economic activity in space that we should compare with the data? What are the main statistics across industries that can inform us on spatial-dynamic linkages? Essentially, we need a tighter connection with the data that goes beyond reduced-form regressions like the ones in Section 5. These are major challenges for the next 50 years of regional science!

REFERENCES

- Baldwin, R.E. and P. Martin. 2004. "Agglomeration and Regional Growth," in J. V. Henderson and J. F. Thisse (eds.), *Handbook of Regional and Urban Economics*, vol. 4. Amsterdam, The Netherlands: Elsevier, pp. 2671–2711.
- Baldwin, R., P. Martin, and G. Ottaviano. 2001. "Global Income Divergence, Trade, and Industrialization: The Geography of Growth Take-Offs," *Journal of Economic Growth*, 6, 5–37.
- Black, D. and J.V. Henderson. 1999. "A Theory of Urban Growth," *Journal of Political Economy*, 107, 252–284.
- Boucekkine, R., C. Camacho, and B. Zou. 2009. "Bridging the Gap Between Growth Theory and the New Economic Geography: The Spatial Ramsey Model," *Macroeconomic Dynamics*, 13, 20–45.
- Brito, P. 2004. "The Dynamics of Growth and Distribution in a Spatially Heterogeneous World," Working Papers, Department of Economics, ISEG, WP13/2004/DE/UECE.
- Brock, W. and A. Xepapadeas. 2008a. "Diffusion-induced Instability and Pattern Formation in Infinite Horizon Recursive Optimal Control," *Journal of Economic Dynamics and Control*, 32, 2745–2787.
- . 2008b. "General Pattern Formation in Recursive Dynamical Systems Models in Economics," MPRA Paper 12305, University of Munich.
- Carlino, G.A., S. Chatterjee, and R. Hunt. 2007. "Urban Density and the Rate of Invention," *Journal of Urban Economics*, 61, 389–419.
- Córdoba, J.C. 2008. "On the Distribution of City Sizes," *Journal of Urban Economics*, 63, 177–197.
- Desmet, K. and M. Fafchamps. 2006. "Employment Concentration across U.S. Counties," *Regional Science and Urban Economics*, 36, 482–509.
- Desmet, K. and E. Rossi-Hansberg. 2009a. "Spatial Growth and Industry Age," *Journal of Economic Theory*, forthcoming.
- . 2009b. "Spatial Development," NBER Working Paper #15349.
- Dumais, G., G. Ellison, and E.L. Glaeser. 2002. "Geographic Concentration as a Dynamic Process," *Review of Economics and Statistics*, 84, 193–204.
- Duranton, G. 2007. "Urban Evolutions: The Fast, the Slow, and the Still," *American Economic Review*, 97, 197–221.
- Duranton, G. and H. Overman. 2005. "Testing for Localization Using Micro-Geographic Data," *Review of Economic Studies*, 72, 1077–1106.
- . 2008. "Exploring the Detailed Location Patterns of U.K. Manufacturing Industries Using Micro-Geographic Data," *Journal of Regional Science*, 48, 213–243.
- Eaton, J. and Z. Eckstein. 1997. "Cities and Growth: Theory and evidence from France and Japan," *Regional Science and Urban Economics*, 27, 443–474.
- Eaton, J. and S. Kortum. 1999. "International Technology Diffusion: Theory and Measurement," *International Economic Review*, 40, 537–570.
- . 2002. "Technology, Geography, and Trade," *Econometrica*, 70, 1741–1779.
- Eeckhout, J. 2004. "Gibrat's Law for (All) Cities," *American Economic Review*, 94, 1429–1451.
- Ellison, G. and E.L. Glaeser. 1997. "Geographic Concentration in U.S. Manufacturing Industries: A Dartboard Approach," *Journal of Political Economy*, 105, 889–927.
- Fujita, M. and J.-F. Thisse. 2002. *Economics of Agglomeration: Cities, Industrial Location, and Regional Growth*, 1st edition. Cambridge, UK: Cambridge University Press, pp. 388–432.
- . 2003. "Does Geographical Agglomeration Foster Economic Growth? And Who Gains and Loses from It?," *Japanese Economic Review*, 54, 121–145.
- Fujita, M., P. Krugman, and A. Venables. 2001. *The Spatial Economy: Cities, Regions, and International Trade*, 1st edition. Cambridge, MA: The MIT Press, pp. 309–328.
- Gabaix, X. 1999a. "Zipf's Law for Cities: An Explanation," *Quarterly Journal of Economics*, 114, 739–767.
- . 1999b. "Zipf's Law and the Growth of Cities," *American Economic Review*, 89, 129–132.
- Griliches, Z. 1957. "Hybrid Corn: An Exploration in the Economics of Technological Change," *Econometrica*, 25, 501–522.

DESMET AND ROSSI-HANSBERG: ON SPATIAL DYNAMICS

- Grossman, G. and E. Helpman. 1991a. *Innovation and Growth in the Global Economy*. Cambridge, MA: The MIT Press.
- . 1991b. "Quality Ladders and Product Cycles," *Quarterly Journal of Economics*, 106, 557–86.
- Hägerstrand, T. 1967. *Innovation Diffusion as a Spatial Process*. Chicago: University of Chicago Press.
- Hillberry, R. and D. Hummels. 2008. "Trade Responses to Geographic Frictions: A Decomposition Using Micro-data," *European Economic Review*, 52, 527–550.
- Holmes, T.J. 2008. "The Diffusion of Wal-Mart and Economics of Density." NBER Working Paper #13783.
- . 2009. "Structural, Experimentalist, and Descriptive Approaches to Empirical Work in Regional Economics," *Journal of Regional Science*, forthcoming.
- Holmes, T.J. and S. Lee. 2008. "Cities as Six-By-Six-Mile Squares: Zipf's Law?," in E. Glaeser (ed.), *The Economics of Agglomeration*. Cambridge, MA: NBER, forthcoming.
- Irwin, E.G. 2009. "New Directions for Urban Economic Models of Land Use Change: Incorporating Spatial Heterogeneity and Transitional Dynamics," *Journal of Regional Science*, forthcoming.
- Jaffe, A., M. Trajtenberg, and R. Henderson. 1993. "Geographic Localization of Knowledge Spillovers as Evidenced by Patent Citations," *Quarterly Journal of Economics*, 108, 577–598.
- Krugman, P. 1991. "Increasing Returns and Economic Geography," *Journal of Political Economy*, 99, 483–499.
- . 1997. *Development, Geography, and Economic Theory*, 1st edition. Cambridge, MA: The MIT Press.
- Lucas, R.E., Jr. 1988. "On the Mechanics of Economic Development," *Journal of Monetary Economics*, 22, 3–42.
- Martin, P. and G. Ottaviano. 1999. "Growing Locations: Industry Location in a Model of Endogenous Growth," *European Economic Review*, 43, 281–302.
- . 2001. "Growth and Agglomeration," *International Economic Review*, 42, 947–68.
- Quah, D. 2002. "Spatial Agglomeration Dynamics," *American Economic Review*, 92, 247–252.
- Rosenthal, S.S. and W.C. Strange. 2008. "The Attenuation of Human Capital Spillovers," *Journal of Urban Economics*, 64, 373–389.
- Rossi-Hansberg, E. 2005. "A Spatial Theory of Trade," *American Economic Review*, 95, 1464–1491.
- Rossi-Hansberg, E. and M. Wright. 2007. "Urban Structure and Growth," *Review of Economic Studies*, 74, 597–624.